

# Regression Odds and Ends

Econ 2560, Fall 2023

Prof. Josh Abel

(Chapters 11.1-2, 8.1-2, 7.2-3)

# Introduction

- We will wrap up the first major section of the course by looking at a few additional fancy things one can do with a regression
  - Binary outcomes (i.e. indicator variable on the lefthand side)
  - Logarithmic variables (i.e. looking at 1% change rather than 1-unit change)
  - Higher-order polynomial specifications
  - Hypothesis tests involving multiple coefficients
  - Testing multiple hypotheses simultaneously

## Binary outcome

- Suppose that instead of earnings, we were interested in the effect of education on *likelihood of earning more than \$20/hr*

$$Y_i = \begin{cases} 1 & \text{if Earnings}_i > 20 \\ 0 & \text{if Earnings}_i \leq 20 \end{cases}$$

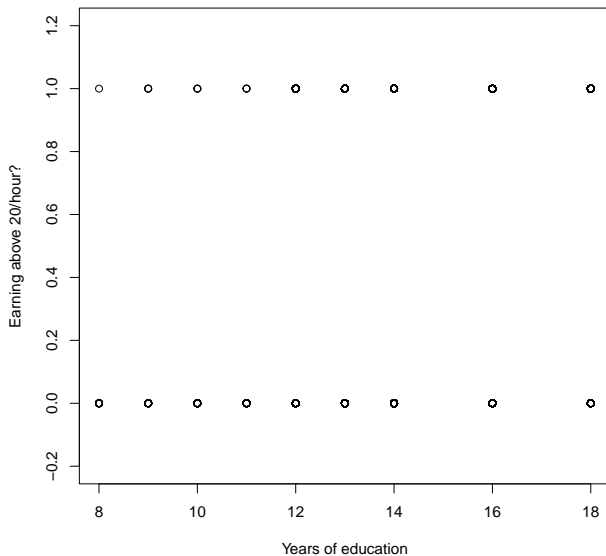
$$E[Y_i | \text{Educ}_i] = \beta_0 + \beta_1 \cdot \text{Educ}_i$$

$$\Pr(Y_i = 1 | \text{Educ}_i) =$$

$$\Pr(\text{Earnings}_i > 20 | \text{Educ}_i) =$$

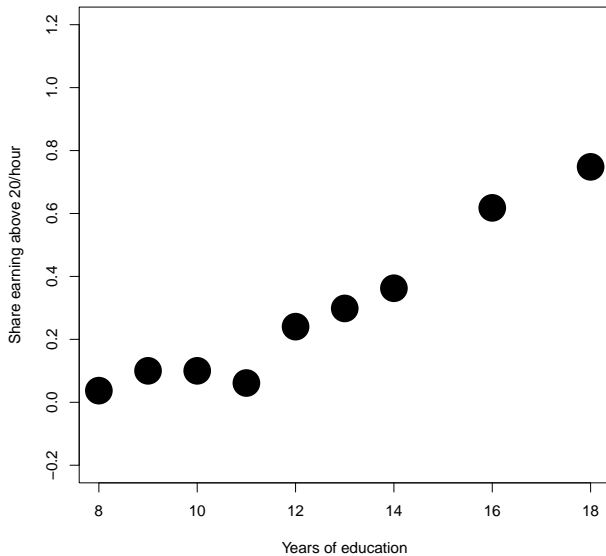
# Binary outcome, visualized

Earning above 20/hour by education



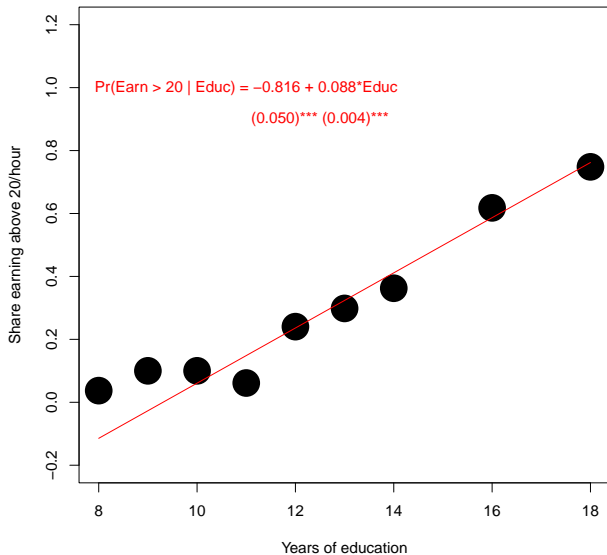
# Binary outcome, visualized

Earning above 20/hour by education



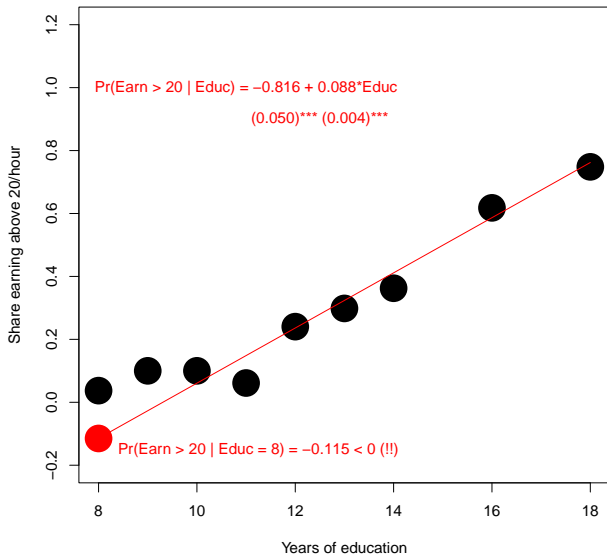
# Binary outcome, visualized

Earning above 20/hour by education



# Binary outcome, visualized

Earning above 20/hour by education



# Linear Probability Model

- Running OLS with a binary outcome is called a “Linear Probability Model” (LPM)
  - Assumes the probability of the outcome is a linear function of the input
- Taken literally, LPM can go haywire
  - May get fitted values outside of  $[0, 1]$ , which are impossible
- Can be addressed by estimating a non-linear function bounded between 0 and 1



# Logit and Probit

- Logit model:

$$\Pr(\text{Earnings}_i > 20 | \text{Educ}_i) = \frac{1}{1 + e^{-(\beta_0 + \beta_1 \cdot \text{Educ}_i)}}$$

- Probit model:

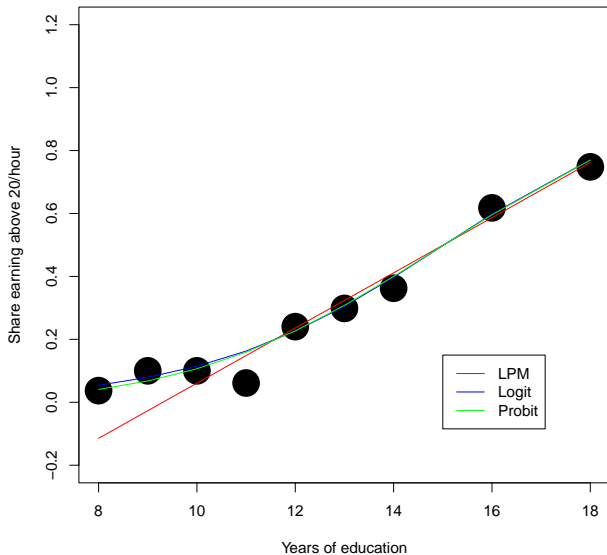
$$\Pr(\text{Earnings}_i > 20 | \text{Educ}_i) = \Phi(\beta_0 + \beta_1 \cdot \text{Educ}_i),$$

where  $\Phi(\cdot)$  is the CDF of the Standard Normal distribution.

- For both Logit and Probit:
  - As  $\beta_0 + \beta_1 \cdot \text{Educ}_i \rightarrow -\infty$ ,  $Pr \rightarrow 0$
  - As  $\beta_0 + \beta_1 \cdot \text{Educ}_i \rightarrow \infty$ ,  $Pr \rightarrow 1$

# Visualizing Logit and Probit

Earning above 20/hour by education



## Outcome: Hourly Earnings Above 20?

	LPM	Logit	Probit
$\hat{\beta}_0$	-0.816 (0.050)***	-6.117 (0.0298)***	-3.724 (0.175)***
$\hat{\beta}_1$	0.088 (0.004)***	0.407 (0.021)***	0.248 (0.012)***

- An additional year of education increases the likelihood of earning above 20 by:
  - LPM: 8.8pp
  - Logit: ??
  - Probit: ??
- Because Logit and Probit are non-linear, there is a different marginal impact at all values of Educ.
  - Can compute marginal effect at specific points of interest (e.g. average)
  - Can compute marginal effect for each observation and then average

# Logit/Probit vs. LPM

- Which was better: LPM or Logit/Probit?

# Logit/Probit vs. LPM

- Which was better: LPM or Logit/Probit?
- Arguments for Logit/Probit:
  - Always gave reasonable predictions (i.e. between 0 and 1)
  - Did a better job of fitting data with low education
- Arguments for LPM:
  - Easy to interpret coefficient (constant marginal effect)
  - Predictions highly similar at medium/high education
- Big picture
  - For pure prediction purposes, Logit/Probit are usually preferable
  - If you want to interpret the marginal effects, LPM is more direct
  - If relationship looks roughly linear, LPM is probably best, even if imperfect
  - If you have enough data, nonparametric regression likely beats them all

## Technical: Estimating Logit and Probit

- Probit (and Logit) conditional means are not linear functions of  $\beta$ s: OLS is no longer applicable
- Could estimate via **Nonlinear Least Squares**: i.e. choose  $\hat{\beta}$ s to minimize

$$\sum_{i=1}^n (Y_i - \Phi(\beta_0 + \beta_1 \cdot \text{Educ}_i))^2 =$$
$$\sum_{Y_i=1} (1 - \Phi(\beta_0 + \beta_1 \cdot \text{Educ}_i))^2 + \sum_{Y_i=0} (\Phi(\beta_0 + \beta_1 \cdot \text{Educ}_i))^2$$

- Same idea as OLS
- But it turns out there is a more efficient (i.e. lower-variance) estimator...

# Maximum Likelihood Estimator (MLE)

- If our Probit model is right, the probability of an observation having  $Y_i = 1$  is:

$$\Phi(\beta_0 + \beta_1 \cdot \text{Educ}_i)$$

- The probability of an observation having  $Y_i = 0$  is:

$$1 - \Phi(\beta_0 + \beta_1 \cdot \text{Educ}_i)$$

- So the probability of us observing our exact dataset is:

$$\prod_{Y_i=1} \Phi(\beta_0 + \beta_1 \cdot \text{Educ}_i) \times \prod_{Y_i=0} (1 - \Phi(\beta_0 + \beta_1 \cdot \text{Educ}_i))$$

- MLE chooses  $\hat{\beta}_0$  and  $\hat{\beta}_1$  to maximize the above expression
  - This is an ugly numerical process, but any statistical package can do it easily, and in fact MLE is softwares' default for models like Logit and Probit

## Marginal effects as percents

- Standard linear model tells us how much \$-earnings rise with 1 additional year of education

$$E[\text{Earnings}_i | \text{Educ}_i] = \beta_0 + \beta_1 \cdot \text{Educ}_i$$

$$\frac{\partial E[\text{Earnings}]}{\partial \text{Educ}} = \beta_1 = \$2.4/\text{hour}$$

- But we may think the more appropriate question is, “What is the *percent change* in earnings from an increase in education?”



# Log specification

- Consider the model below:

$$E[\ln(\text{Earnings})_i | \text{Educ}_i] = \beta_0 + \beta_1 \cdot \text{Educ}_i$$

- Interpretation of  $\beta_1$  changes

$$\begin{aligned}\beta_1 &= \frac{\partial E[\ln(\text{Earnings})]}{\partial \text{Educ}} \\ &= \frac{1}{\text{Earnings}} \cdot \frac{\partial E[\text{Earnings}]}{\partial \text{Educ}} \\ &= \frac{\partial E[\text{Earnings}]/\text{Earnings}}{\partial \text{Educ}}\end{aligned}$$

- So  $\beta_1$  gives the percent change in earnings from 1 additional year of education

## Log specification results

Y	X	$\beta_0$	$\beta_1$	Interpretation
Earn	Educ	-12.1 (1.4)***	2.4 (0.1)***	Educ $\uparrow$ 1y $\rightarrow$ Earn $\uparrow$ \$2.4/hr
ln(Earn)	Educ	1.37 (0.06)***	0.11 (0.00)***	Educ $\uparrow$ 1yr $\rightarrow$ Earn $\uparrow$ 11%
Earn	ln(Educ)	-63.2 (3.6)***	32.1 (1.4)***	Educ $\uparrow$ 1% $\rightarrow$ Earn $\uparrow$ \$0.32/hr
ln(Earn)	ln(Educ)	-1.03 (0.15)***	1.50 (0.06)***	Educ $\uparrow$ 1% $\rightarrow$ Earn $\uparrow$ 1.5%

- Note that the final specification (log-log) has the special interpretation of an **elasticity** and is commonly used

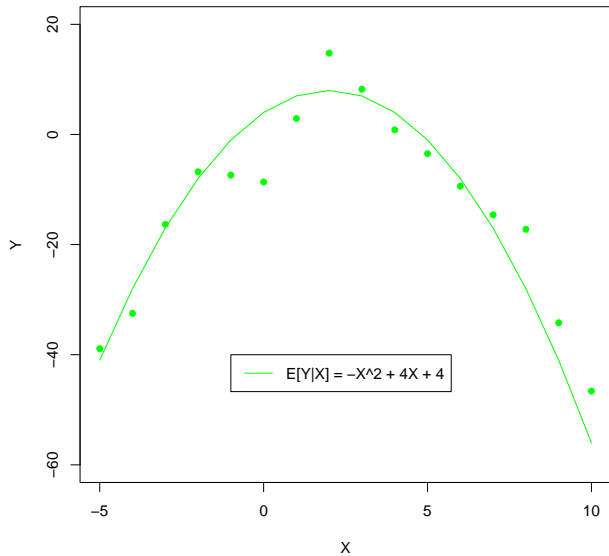
# A quadratic regression function

- Suppose we suspected the true regression function was actually quadratic

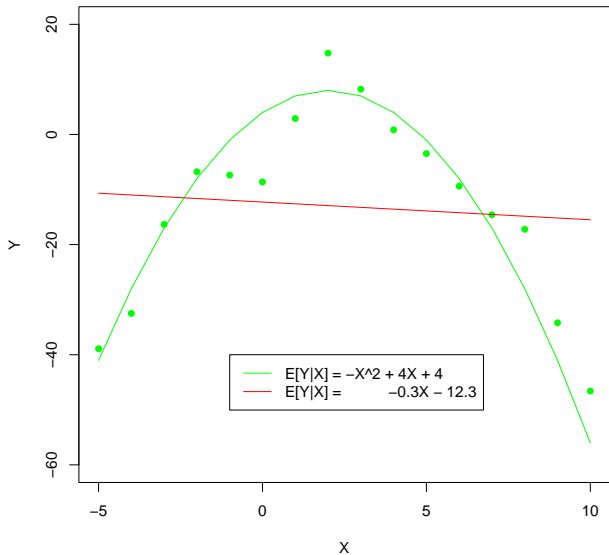
$$E[Y_i|X_i] = \beta_0 + \beta_1 \cdot X_i + \beta_2 \cdot X_i^2$$

- This can be a critical adjustment if the relationship is non-monotonic
  - A line will never usefully summarize relationship that is increasing in one region and decreasing in another
  - E.g. Impact of caloric intake on health outcomes

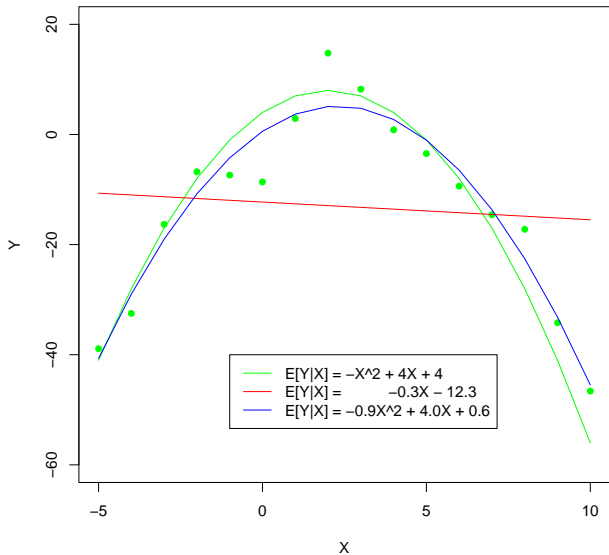
# A quadratic regression function



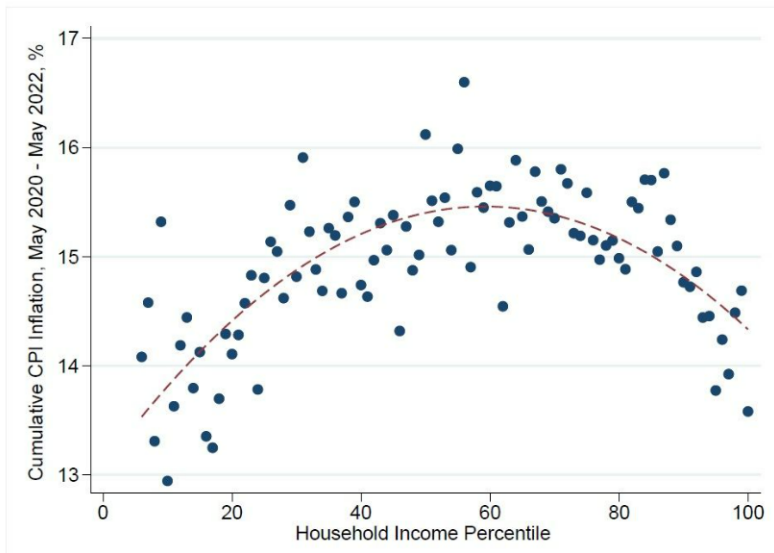
# A quadratic regression function



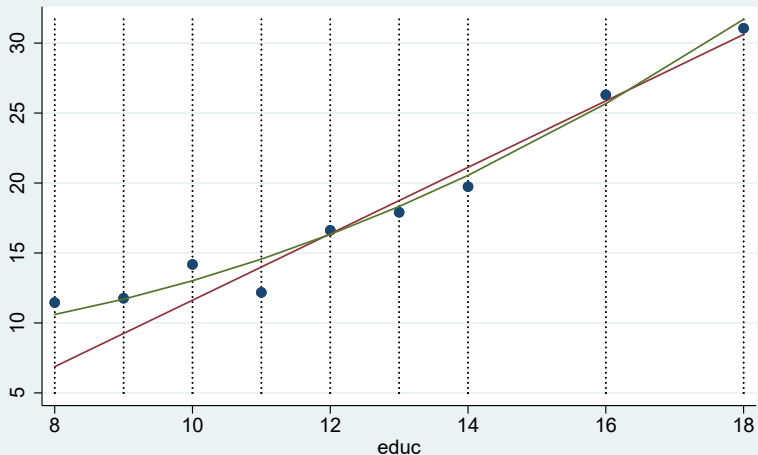
# The quadratic regression estimator



## Inflation Inequality by Income Percentile



# Back to Earnings...





# Regression Output

	(1)	(2)
Education	2.37	-0.83
(SE)	(0.10)***	(0.90)
Education <sup>2</sup>		0.11
		(0.03)***
Constant	-12.12	10.00
	(1.36)***	(5.97)*

Marginal effect of education is  $-0.83 + 2 \cdot 0.11 \cdot \text{Education}$ .

# “How do I choose the polynomial?”

- Look at the data!
  - Scatterplot
  - If helpful, “binned scatterplot” / non-parametric approach
- If monotonic and broadly linear, linear should be fine
  - Still, probably worth checking a quadratic
- If clearly non-linear or – especially – non-monotonic, use quadratic
  - Worth checking a cubic term
- If relationship changes direction *multiple times* or otherwise looks wacky, can try higher-order polynomials
  - At that point, strongly consider non-parametric approach
    - High-order polynomials are hard to interpret anyway (unlike linear and quadratic), so not much advantage over non-parametric

## Linear combination of coefficients

$$\hat{E}[Y_i|X_i] = 10.00 - 0.83 \cdot \text{Educ}_i + 0.11 \cdot \text{Educ}_i^2$$

(5.97)*	(0.90)	(0.03)***
$\hat{\beta}_0$	$\hat{\beta}_1$	$\hat{\beta}_2$

- Suppose we want to know the marginal impact of education when  $\text{Educ} = 10$ 
  - $\hat{\beta}_1 + 2 \cdot \hat{\beta}_2 \cdot 10 = -0.83 + 2 \cdot 0.11 \cdot 10 = 1.37$
  - At 10 years of education, the marginal effect of education is \$1.37/hr.
- What is the Standard Error of  $\hat{\beta}_1 + 2 \cdot \hat{\beta}_2 \cdot 10$ ?
  - $SE(\hat{\beta}_1 + 20 \cdot \hat{\beta}_2) = SE(\hat{\beta}_1) + 20 \cdot SE(\hat{\beta}_2)$ ?

## Linear combination of coefficients

$$\hat{E}[Y_i|X_i] = 10.00 - 0.83 \cdot \text{Educ}_i + 0.11 \cdot \text{Educ}_i^2$$

(5.97)*	(0.90)	(0.03)***
$\hat{\beta}_0$	$\hat{\beta}_1$	$\hat{\beta}_2$

- Suppose we want to know the marginal impact of education when  $\text{Educ} = 10$ 
  - $\hat{\beta}_1 + 2 \cdot \hat{\beta}_2 \cdot 10 = -0.83 + 2 \cdot 0.11 \cdot 10 = 1.37$
  - At 10 years of education, the marginal effect of education is \$1.37/hr.
- What is the Standard Error of  $\hat{\beta}_1 + 2 \cdot \hat{\beta}_2 \cdot 10$ ?
  - $SE(\hat{\beta}_1 + 20 \cdot \hat{\beta}_2) = SE(\hat{\beta}_1) + 20 \cdot SE(\hat{\beta}_2)$ ?
  - No.
  - $SE(\hat{\beta}_1 + 20 \cdot \hat{\beta}_2) = \sqrt{\text{var}(\hat{\beta}_1) + 20^2 \cdot \text{var}(\hat{\beta}_2) + 2 \cdot 20 \cdot \text{cov}(\hat{\beta}_1, \hat{\beta}_2)}$

# Variance-covariance matrix

- OLS estimates do not just generate a variance (or SE) for each coefficient...
- ... but also covariances between all pairs of coefficients!
- These are organized into a **“variance-covariance matrix”**

$$\begin{bmatrix} \hat{\sigma}_0^2 & \hat{\sigma}_{0,1} & \hat{\sigma}_{0,2} \\ \hat{\sigma}_{0,1} & \hat{\sigma}_1^2 & \hat{\sigma}_{1,2} \\ \hat{\sigma}_{0,2} & \hat{\sigma}_{1,2} & \hat{\sigma}_2^2 \end{bmatrix}$$

# Variance-covariance matrix

- OLS estimates do not just generate a variance (or SE) for each coefficient...
- ... but also covariances between all pairs of coefficients!
- These are organized into a **“variance-covariance matrix”**

$$\begin{bmatrix} \hat{\sigma}_0^2 & \hat{\sigma}_{0,1} & \hat{\sigma}_{0,2} \\ \hat{\sigma}_{0,1} & \hat{\sigma}_1^2 & \hat{\sigma}_{1,2} \\ \hat{\sigma}_{0,2} & \hat{\sigma}_{1,2} & \hat{\sigma}_2^2 \end{bmatrix}$$

Square roots give SEs for individual coefficients

# Results: "lincom"

```
. regress ahe educ educ2, robust
```

```
Linear regression              Number of obs   =      2,731
                              F(2, 2728)      =      308.30
                              Prob > F            =      0.0000
                              R-squared           =      0.1873
                              Root MSE        =      11.221
```

	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
ahe						
educ	-.8281024	.8989831	-0.92	0.357	-2.590859	.9346542
educ2	.1129571	.0332398	3.40	0.001	.0477794	.1781347
_cons	10.00201	5.970018	1.68	0.094	-1.704205	21.70822

```
. mat list e(V)
```

```
symmetric e(V) [3,3]
      educ      educ2      _cons
educ   .80817067
educ2  -.02973218  .00110488
_cons  -5.3398927  .19456393  35.641118
```

```
. lincom educ+20*educ2
```

```
( 1)  educ + 20*educ2 = 0
```

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
(1)	1.431039	.2466506	5.80	0.000	.9473977	1.91468

# Results from cubic model

```
. regress ahe educ educ2 educ3, robust
```

Linear regression

```
Number of obs   =    2,731
F(3, 2727)      =    209.99
Prob > F        =    0.0000
R-squared       =    0.1878
Root MSE       =    11.22
```

ahe	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
educ	-8.183991	5.518622	-1.48	0.138	-19.0051	2.637112
educ2	.6703205	.4379484	1.53	0.126	-.1884237	1.529065
educ3	-.0137418	.0112703	-1.22	0.223	-.0358409	.0083574
_cons	41.57251	22.55833	1.84	0.065	-2.660629	85.80565

- No individual coefficient on Education is significant.
- So is Education not a statistically significant determinant of earnings?



# Testing Multiple Hypotheses

- Two statements:
  - ①  $\hat{\beta}_1$  is insignificant,  $\hat{\beta}_2$  is insignificant,  $\hat{\beta}_3$  is insignificant
  - ②  $\hat{\beta}_1, \hat{\beta}_2, \hat{\beta}_3$  are all *jointly* insignificant
- These statements are very different
  - #1 is done one-at-a-time using our standard z-test
  - #2 requires an “F-test”

# F-test on cubic model

```
. regress ahe educ educ2 educ3, robust
```

Linear regression

```
Number of obs   =    2,731
F(3, 2727)      =    209.99
Prob > F        =    0.0000
R-squared       =    0.1878
Root MSE       =    11.22
```

ahe	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
educ	-8.183991	5.518622	-1.48	0.138	-19.0051	2.637112
educ2	.6703205	.4379484	1.53	0.126	-.1884237	1.529065
educ3	-.0137418	.0112703	-1.22	0.223	-.0358409	.0083574
_cons	41.57251	22.55833	1.84	0.065	-2.660629	85.80565

- Model isn't sure whether it's the linear, quadratic, or cubic term that matters...
- But it knows that together ("jointly"), they do matter!

# F-test on cubic model

```
. regress ahe educ educ2 educ3, robust
```

```
Linear regression
```

```
Number of obs   =    2,731  
F(3, 2727)      =    209.99  
Prob > F        =    0.0000  
R-squared       =    0.1878  
Root MSE       =    11.22
```

ahe	Robust		t	P> t	[95% Conf. Interval]	
	Coef.	Std. Err.				
educ	-8.183991	5.518622	-1.48	0.138	-19.0051	2.637112
educ2	.6703205	.4379484	1.53	0.126	-.1884237	1.529065
educ3	-.0137418	.0112703	-1.22	0.223	-.0358409	.0083574
_cons	41.57251	22.55833	1.84	0.065	-2.660629	85.80565

```
. test (educ=0) (educ2=0) (educ3=0)
```

- ( 1) educ = 0
- ( 2) educ2 = 0
- ( 3) educ3 = 0

```
F( 3, 2727) = 209.99  
Prob > F = 0.0000
```

## F test with homoskedasticity

- The joint hypotheses above can be calculated in the following way when  $\text{var}(u_i|X) = \sigma_u^2$ , a constant (i.e. under homoskedasticity)
- Test statistic is:

$$F = \frac{(R_{unr}^2 - R_{res}^2)/q}{(1 - R_{unr}^2)/(n - k - 1)},$$

where  $R_{unr}^2$  is the  $R^2$  from the regression,  $k$  is the number of regressors,  $R_{res}^2$  is the  $R^2$  from a regression assuming  $H_0$  is true,  $q$  is the number of restrictions, and  $n$  is the number of observations

$$F \sim F_{q, n-k-1},$$

# Intuition of F test

- We know  $R_{unr}^2 > R_{res}^2$  because it is unconstrained
- That doesn't mean the restrictions are wrong – could be random noise
- The F test tests whether the reduction in  $R^2$  from imposing  $H_0$  is “statistically significant”
  - The model fits worse, but how much worse?

## Final comments on F tests

- Without homoskedasticity, formulas are much messier, but intuition is similar
- The F test is basically a generalization of a t test
  - t tests simply set  $q = 1$  – i.e. only 1 restriction
- Can also construct “**confidence sets**”